

# Biased Policy and Political Behavior: The Case of Uneven Removal of Elected Mayors in Israel

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Unlike the previous literature on mass policy feedback, the present study argues that a negative message embodied in public policy may foster or dampen political participation depending on social group affiliation. The policy change we use to examine the effect of biased policy (a perceived negative message) on political behavior is the removal of elected mayors that were replaced by an appointed committee in a large number of Arab and Jewish municipalities in Israel, which was skewed significantly toward Arab municipalities. We show that Arab voters in intervened municipalities are more likely to show up at the ballot boxes in national elections and they tend to vote more for Arab parties. In contrast, the political participation of Jewish citizens is lower in municipalities with an appointed council without noticeable effect on vote choice.

**Keywords:** Biased Policy, Negative Message, Political Participation, Political Behavior, Israel, Policy Feedback, Mayoral Politics in the Middle East, Arab and Jewish Municipal Politics, Voter Choice, Social Group Affiliation, National Elections Israel, Internal Political Efficacy, External Political Efficacy.

#### **Related Articles:**

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aboahma711122. 2009. "Video clip by Raam Taal Party." https://www. youtube.com/watch?v=JEzDRjk9RQw

A diferencia de la literatura anterior sobre la retroalimentación de la política de masas, el presente documento argumenta que un mensaje negativo incorporado en la política pública puede fomentar o atenuar la participación política en función de la afiliación del grupo social. El cambio de política que usamos para examinar el efecto de una política sesgada (un mensaje negativo percibido) sobre el comportamiento político es la destitución de alcaldes electos que fueron reemplazados por un comité designado en un gran número de municipalidades árabes y judías en Israel que estaba sesgada significativamente hacia Municipios árabes. Demostramos que los votantes árabes en municipios intervenidos son más propensos a llegar a las urnas de votación en las elecciones nacionales y tienden a votar más por partidos árabes. Por el contrario, la participación política de los ciudadanos judíos es menor en los municipios con un consejo designado sin un efecto notable en la elección del voto.

**Palabras Clave:** Retroalimentación Política, Participación Política, Elección de Voto, Afiliación a un Grupo Social, Elecciones Nacionales, Eficacia Política Interna, Eficacia Política Externa, Políticas y Política.

**摘要:**和以往研究大型政策反馈的文献所不同的是,本文主张,公共政策中的一个负面信息可能会根据社会团体的隶属关系而促进或抑制政治参与。笔者以一种政策变化为例,检验了偏见性政策(即被感知为负面信息的政策)对政治行为产生的效果。该政策变化则是,以色列的许多阿拉伯和犹太市政开除当选市长,并用指派的委员会予以替代(这些委员会大多数都偏向于阿拉伯市政当局)。笔者表明,在市政当局被干预的情况下,阿拉伯投票者更可能出现在国家选举的投票箱前,并且他们倾向于更多地投给阿拉伯党派。相反,犹太公民在由指派委员会领导市政的情况下较少进行政治参与,并且对投票选择没有产生显著作用

**关键词:**政策反馈,政治参与,投票选择,社会团体隶属关系,国家选举, 内在政治效能感,外在政治效能感,政策和政治.

Policies are not only the outcome of politics, they also shape the political process through the messages they convey and they subsequently impact upon

future policies. Unlike the previous literature on mass policy feedback, the present article argues that negative messages embodied in public policy may foster or dampen political participation, depending on social group affiliation, and context. Introducing the interplay between social group consciousness and biased policy associated with a perceived negative message such as discrimination leads to higher, rather than lower, political participation of certain groups.<sup>1</sup> Our conceptual framework postulates that a biased policy provokes social group consciousness regarding the relative position in society among individuals with minority group affiliation, which in turn contributes to a higher sense of internal efficacy and consequently generates the unusual combination of low external political efficacy (EPE) with high internal political efficacy (IPE). As a result, the target population with minority affiliation is expected to be more active in the political process as long as they have at least minimal faith in the political system.

However, the same policy may dampen the political participation among the target population that belongs to the majority group as a result of a lower level of EPE and IPE. In contrast to the minority group, our expectation regarding the target population affiliated with the majority group resembles the political behavior predicted by previous scholars (see e.g., Schneider and Ingram 1993). In this article, we examine this theoretical prediction in a particular context— but it may explain also political behavior in other contexts. We believe that it is particularly relevant to understand the political participation of certain social groups in the United States following the new immigration policy and intense rhetoric that key players use, including the President Trump, with respect to immigrants from certain parts of the world.

The policy change we use to examine the effect of biased policy on political behavior is the removal of elected mayors that were replaced by an appointed committee in a large number of Arab and Jewish municipalities in Israel in the 2000s. The central government intervention was not uniformly distributed across sectors; the appointment of summoned committees was skewed significantly toward Arab municipalities. The policy intervention in a relatively large number of Arab municipalities (more than their share) has been perceived as a discriminatory act by Arab citizens. We show that Arab citizens in intervened municipalities are more likely to show up at the ballot boxes in national elections, and that they tend to vote more for Arab parties as compared to Arab municipalities without intervention. In contrast, the political participation of Jewish citizens is

<sup>1</sup>Adding the nature of interaction between social groups and, in particular, between the dominant group and minority/ethnic groups seems to be one of the more promising routes to account also for the significant variation across various social groups in political engagement and vote choice. The differences in group political engagement have been empirically attributed to group identity or consciousness (see e.g., Ben-Bassat and Dahan 2012; Miller *et al.* 1981; Wilcox and Gomez 1990), ethnic-based institutions that encourage mobilization (see e.g., Brown and Brown 2003; Harris 1994; Rosenstone and Hansen 1993; Tate 1991; Verba, Schlozman, and Brady 1995), and political empowerment (see e.g., Bobo and Gilliam 1990; Logan, Darrah, and Oh 2012; Pantoja and Segura 2003; Tate 1994; Washington 2006).

lower in municipalities with an appointed council without noticeable effect on vote choice as compared to the general Jewish population.

In his influential article, Pierson (1993) shows that public policy may affect politics also by altering individuals' resources and incentives.<sup>2</sup> Yet the policy feedback literature devoted significant research efforts to explore the interpretive effects of public policy on politics as compared to the effects of resources (Verba, Schlozman, and Brady 1995) or incentive channels (Campbell 2003). Schneider and Ingram (1993) offer an intriguing theoretical argument, that certain groups are expected to have a negative experience with government, which dampens their standard modes of political participation such as voting.

In a series of empirical works, scholars have studied the effect that personal experience with public policy may have on conventional forms of political behavior in an attempt to uncover the link between policies and politics. Soss (1999) was one of the first to find that recipients of means-tested welfare programs are less active than recipients of Social Security Disability Insurance, even after controlling for individual characteristics. Using in-depth interviews, Soss attributed the lower level of political participation to the negative experience that clients of means-tested programs went through, such as ongoing scrutiny to prove their eligibility, threats of termination of welfare benefits, and the compulsion to share intimate information. Bruch, Marx Ferree, and Soss (2010) show that political participation varies even within means-tested programs. They employ a convincing identification strategy to examine three types of means-tested programs of public assistance (Head Start, Housing Aid, and Temporary Assistance to Needy Families), and find that the recipients of a welfare program with a paternalistic design are less likely to participate in the political process. As expected, the positive policy experience of veterans benefitting from the GI Bill has fostered their political and civic participation (Mettler 2005). These important studies, among others, lend strong empirical support to the claim that negative interpretive messages reduce political participation and positive experiences enhance it.<sup>3</sup>

Miller and Krosnick (2004) have already examined the effects of perceived undesirable policy change on political behavior. They exploited a field experiment and showed that a letter highlighting undesirable policy changes, which was sent to potential contributors, increased the number of financial contributions made to a political lobbying organization. However, this interesting research has not studied that effect in the social context of ethnic and minority groups. Other scholars have investigated the importance of social context, but not in relation to policy making. Schildkraut (2005) found that individual-perceived discrimination (self-reported) has a positive effect on political engagement of Latinos. In addition, several studies examine the effects of general political climate on political

 $<sup>^2</sup>$  The potential effect of policy on politics was first explored in the seminal work of Schattschneider (1935) on the politics of tariffs.

<sup>&</sup>lt;sup>3</sup> Campbell (2012) provides a review of this literature with reference to the effects of policy in other fields on politics. Note that Soss and Schram (2007) do not find evidence of welfare policy reforms changing public opinion.

participation of ethnic and minority groups. Pantoja, Ramirez, and Segura (2001) find that Latinos who chose to naturalize in California between 1992 and 1996 are more likely to vote, as compared to Latinos who had naturalized in California before that time period and as compared to their contemporaries in Texas and Florida. They attributed this effect to the hostile political environment toward Latinos in California in that period of time. Cho, Gimpel, and Wu (2006) have also offered evidence on the effect of hostile environments on political participation. They have shown that the political participation of Arab Americans has increased after 9/11, as a result of a hostile climate, which was measured by the prevalence of the terms "Patriot Act" or "War in Iraq" in the news. More recently, White (2016) found that immigration enforcement in the United States led to an increase in Latinos voter turnout by two to three percentage points. Interestingly, Latino citizens were mobilized by a policy change that did not target them personally.

This article advances our understanding of the interpretive effects embodied in public policy on political behavior. A perceived negative message may produce the opposite political response from the one that was traditionally suggested in the mass policy feedback literature. Employing a general policy change allows us to provide empirical support for that prediction by comparing the political behavior of the target population and general population in both majority and minority groups following that policy change.

Our second contribution refers to the role of public policy in explaining the variation in political behavior across social strata, and in particular that of subordinate-minority groups, which could not be accounted for by differences in individual or institutional characteristics. While the previous literature has made an important contribution in studying the individual and institutional factors, our understanding of political behavior of ethnic and minority groups is still at its infant stage.<sup>4</sup> It is imperative to study the differences in political behavior of social groups in light of the changing social structure of nation states that is far more heterogeneous today than it was after World War II.

A third contribution of our empirical examination is in providing an explanation that may account for the substantial fluctuations over time in political participation and political preferences. Standard prevailing explanations of political participation and vote choice that rest on relatively stable resources, such as education or income and mobilization strategies, could not explain the relatively large fluctuations in voter turnout from election to election.

In the next section, we describe the centralization program that was launched in Israel in the 2000s and the attitudes toward that policy change among different social groups. We then lay out a conceptual framework that guides our empirical investigation regarding the link between biased public policy and

<sup>&</sup>lt;sup>4</sup>Recently, Logan, Darrah, and Oh (2012), as well as Leighley and Nagler (2013), have shown once again that certain groups, such as African Americans, vote at higher rates than Anglos, while the political engagement of other groups (Latinos and Asian Americans) is lower as compared with Anglos, after controlling for standard socioeconomic status factors such as education, income, and age. See also Tate (1991, 1994) and Leighley (2001) for similar findings.

political participation and voting behavior of minority and majority groups. The subsequent section presents the empirical analysis before concluding.

#### **Policy Change and Attitudes**

This research focuses on the potential difference in political reaction to a major policy change in terms of political participation and political preferences of citizens that belong to dominant and minority groups. The same policy change may be perceived differently by individuals, conditional on their social group affiliation, and regardless of the true motivations that stand behind that policy act.

#### **Policy Change**

The policy change we use is the appointment of a summoned committee in a large number of Arab and Jewish municipalities in Israel. Following a severe fiscal crisis in municipalities in Israel during 2001-03, the central government launched a major economic program that affected more than half of the Jewish municipalities and more than 90 percent of Arab municipalities (Ben-Bassat and Dahan 2009). To cope with that fiscal crisis, the Interior Ministry employed two "soft" measures of administrative subordination (an appointment of an accountant who has to approve in advance every expense of the municipality and the imposition of a recovery program) and a more intrusive measure which removes the elected mayor and members of the local council and appoints a summoned board as the sole authority of the municipality until the next local election.<sup>5</sup> According to stated policy, the interior ministry removed elected officials who were replaced by appointees in municipalities with a budget deficit over 15 percent of its total revenue, short-term debt over 30 percent of its total revenue, and tax effectiveness indicator below a certain threshold. In addition, the elected local council was dismissed if it did not approve the budget within three months of the beginning of the fiscal year.

In a relatively short period of time, from 2005 through 2011, the elected mayors and members of local councils were replaced by a summoned board in 30 municipalities. Indeed, the number of new summoned boards in the last decade exceeds that of all previous five decades combined (Ben-Bassat, Dahan, and Klor 2016b). This policy change was not uniformly distributed: a summoned committee was the sole authority in 21 out of 80 Arab municipalities as compared to 9 out of 110 Jewish municipalities.<sup>6</sup> This is in contrast with the former

<sup>&</sup>lt;sup>5</sup> See Ben-Bassat, Dahan, and Klor (2016a), for a detailed description of the local election system in Israel.

<sup>&</sup>lt;sup>6</sup> A municipality is defined Jewish (Arab) if more than a half of its residents are of Jewish (Arab) origin. In general, Jews and Arabs reside in separate municipalities except eight large cities like Jerusalem and Tel-Aviv where both Jews and Arabs live together (the average share of Arabs in those cities is around 20 percent). We exclude regional municipalities due to data limitations.

actual policy of the central government where the very few summoned committees that had been appointed in the past were distributed evenly between Jewish and Arab municipalities (i.e., according to their shares in the population).

The implementation of the policy to appoint summoned committees was determined according to prespecified rules and with very little discretion, thus leaving slim room for discrimination. Ben-Bassat, Dahan, and Klor (2016b) have shown that the imposition of appointed committees was closely related to local fiscal stress measures. Yet that policy reform may still be perceived as discriminatory by Arab residents. The poor tax capacity of many Arab municipalities may reflect general policy biases, such as government refusal to expand the jurisdiction area of Arab municipalities, lack of commercial and industrial zones in Arab municipalities (due to past biased incentive policies), and equalization grant policy. Arab leaders frequently claim that the poor fiscal conditions in Arab municipalities are mainly the result of those disparities that are rooted in past and present discrimination. Therefore, while the implementation of the rules that govern the appointment of summoned committees might not be ethnically biased, Arab citizens may still perceive that policy as discriminatory.

#### Attitudes toward Appointed Committees by Ethnic Origins

While the central government has stated that poor local management is the key factor behind fiscal crises, others blame the combined effect of general economic slowdown (that leads to lower revenue) and the dramatic cuts in equalization grants to poor municipalities as the main causes of fiscal distress in the beginning of the 2000s. As a result, one could blame either the central government or local elected officials for the ballooned debts and deficits of municipalities. The uncertainty regarding the root cause of local fiscal crises opens the door to different interpretations of central government intervention. The same policy may be perceived differently by various social groups depending, in part, on the level of trust that individuals have in the government. According to the Israel Democracy Institute's yearly survey, Arab citizens in Israel consistently show a lower degree of trust in government as compared to Jewish citizens. Around 60 percent of Arab respondents do not trust the government at all, relative to 33 percent among Jewish individuals (Hermann *et al.* 2013, 42).

Beeri and Yuval (2013) provide direct evidence on the level of support for central government intervention in running local municipalities by ethnic origin (i.e., national majority and minority groups). Their survey of 1,321 residents in 156 municipalities specifically covers questions on the respondents' attitudes toward appointed committees. For example, one of the questions is: "In cases of local crisis, replacing the elected leadership with a convened committee appointed by the Ministry of Interior is an appropriate decision" (Beeri and Yuval 2013, 641). They found that residents of municipalities without an appointed committee tend to support such intervention, but residents in intervened municipalities

are more likely to oppose it. They show also that Arab residents are more likely to oppose the removal of elected local mayors and councilors as compared to Jewish residents. Taken together these two findings imply that the attitudes of residents in municipalities with an appointed committee are different conditional on their ethnic (minority) origin.<sup>7</sup>

Arab citizens in municipalities view the act of removing their elected mayors and council members and replacing them by the Interior Ministry appointees as a threat to their fundamental rights rather than a tool to guarantee adequate level of local services (El-Taji 2008). The dissatisfaction of Arab citizens has translated into frequent violent instances against the Interior Ministry appointees, including two failed attempts to assassinate the head of the summoned committee in Tayibe and Turan. The national Arab leaders in the Israeli parliament have headed the opposition to these policy measures and, in particular, to summoned committees. In fact, Arab parties in the parliament have issued a motion of no confidence against the government in 2009 to express their resistance to appointed local councils in Arab municipalities. One might find signs of dissatisfaction also in Jewish municipalities regarding summoned committees, but there was no national party that challenged the government on this particular issue and, naturally, discrimination was not brought up.

As mentioned before, the appointed council has been used disproportionally in Arab municipalities and against the will of their residents. Moreover, all of the summoned committees in Arab municipalities were first headed by appointees of Jewish origin. These two facts have been used by national and local Arab leaders in Israel to claim that this policy change is discriminatory in nature. The National Committee of Chairmen of Arab Local Authorities (NCCAL), which is a leading organization of Arab citizens in Israel, has expressed the strongest opposition to the use of appointed committees in Arab municipalities and has organized a general strike in all Arab municipalities following the decision of the Interior Ministry to dismiss local elected officials.<sup>8</sup> The spokesmen of NCCAL also voiced this dissatisfaction in the international outlet of Kul el Arab, stating that the policy tool of an appointed committee targets the Arab population and undermines the legitimacy of its elected leadership.<sup>9</sup> The word "discrimination" may be found numerous times in Arab NGO reports or websites and in the popular press in Israel regarding summoned boards.<sup>10</sup> Naturally, this type of claim has not been raised in Jewish municipalities with a summoned committee. Thus there is much evidence to conclude that this same policy change

<sup>&</sup>lt;sup>7</sup> In the multiple regression analysis, Beeri and Yuval (2013) found that ethnicity (minority) has a significant coefficient while the appointed committee has the expected sign but is insignificant. However, the coefficient of an appointed committee becomes significant after excluding ethnicity (minority).

<sup>&</sup>lt;sup>8</sup> See the *NRG* website, on June 29, 2007.

<sup>&</sup>lt;sup>9</sup> See *Kul al Arab*, July 6, 2007 (p. 4).

<sup>&</sup>lt;sup>10</sup> See for example the report for 2012 published by The Coalition Against Racism in Israel (p. 24). See also the column by Ali Hider in *TheMarker (Haaretz)* on October 18, 2012.

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has been perceived differently by minority and dominant groups. In particular, the central government intervention led to an increased sense of threat and dissatisfaction especially among Arab residents. In the empirical analysis, we use the appointed summoned committee that was instituted by the central government as our (indirect) measure of dissatisfaction or perceived discrimination.

#### Modeling the Link between Public Policy and Political Behavior

We argue here that public policy is linked to political behavior through its differential impact on EPE (defined as the intensity of beliefs that the government is responsive) and IPE (defined as the feeling that an individual political action could have an impact on the political process). The main core of our thesis is that the removal of local elected mayors and councilors produces an unusual mixture of lower EPE and higher IPE among the target population (as compared to the general population) in the minority group, that fosters their political participation. In contrast, the same policy generates a standard combination of lower external and IPE among the target population (as compared to the general population) in the majority group that depresses their political participation. Fluctuations in the level of political efficacy as a result of a biased policy change may translate into changes in political participation and vote choices. This prediction rests on prior research that has established a solid link between both dimensions of political efficacy and political participation (Finkel 1985; Niemi, Craig, and Mattei 1991; Rosenstone and Hansen 1993; Shingles 1981; Verba and Nie 1972).<sup>11</sup>

We argue that ethnically biased public policy as perceived by the minority group raises group consciousness, and following Shingles (1981), that contributes to a higher level of IPE among individuals of the target population who are affiliated with that minority group. We borrow the concept of group consciousness from Miller and others (1981) who argue that "group consciousness involves identification with the group and a political awareness or ideology regarding the group's relative position in society along with a commitment to collective action aimed at realizing the group interest."<sup>12</sup> The last part of this definition of group consciousness-based IPE.

As we discuss above, there is strong evidence that Arab citizens—especially in intervened municipalities—view the imposition of appointed committees as an act of discrimination by the central government. As Beeri and Yuval (2013)

<sup>&</sup>lt;sup>11</sup> See Morrell (2003) for extensive review of this line of research.

<sup>&</sup>lt;sup>12</sup> The definition of Miller and others (1981, 495) of group consciousness covers four components: social identification, the extent of in-group positive affect and out-group dislike (*polar affect*), the level of satisfaction with the group's relative position (*polar power*), and the perceived responsibility of the political system for the group's relative position in society (*individual vs. system blame*). Our conceptual framework emphasizes the likely interaction between public policy and the effect of the last component of group consciousness definition on IPE and EPE.

show, the reduction in local autonomy of Arab residents has been done by the central government against the will of citizens in intervened Arab municipalities (the Arab target population). We assume that the policy intervention raises group consciousness of Arab citizens and political awareness regarding their deprived position in the general population and their dedication to political action to promote their cause. As suggested by Shingles (1981), rising group consciousness leads to a higher sense of IPE among Arab citizens in intervened municipalities. Consequently, a higher level of political participation on behalf of the group and more votes to affiliated parties are natural actions to advance their interests.

A large body of studies has established that various aspects of group membership such as instrumental motives (see e.g., Uhlaner 1989), social identity (see e.g., Tajfel 1981), social pressure (see e.g., Grossman and Helpman 2001), and group-rule-utilitarian approach à la Harsani (1980), may be important in explaining individual's voting behavior in large elections.<sup>13</sup> The predictions of these theories that individuals will do their part to help the group, have gained empirical support in both empirical and laboratories studies. For example, Uhlaner (1989) found that individuals who are associated with a labor union were likely to vote than nonunion individuals. More recently, employing last name as a measure of group affiliation, Ben-Bassat and Dahan (2012) presented evidence that individuals in Arab municipalities in Israel were more likely to vote for a candidate who shares their last name (social group) than to other candidates. Using laboratory experimentation methods, Großer and Schram (2006) and Fowler and Kam (2007) have also shown positive relations between group membership and voter turnout.

Unlike Jewish municipalities, Arab municipalities already had the organizations needed to exploit the increased sense of dissatisfaction that is associated with the policy change to gain more votes. A body of evidence suggests that political mobilization is a direct response to the degree of threat and discrimination a group experiences (Austin, Middleton, and Yon 2012; Campbell 2003; Feldman and Stenner 1997; Marcus, Neuman, and Mackuen 2000; Miller *et al.* 1981; Rudolph, Gangl, and Stevens 2000; Salamon and Van Evera 1973; Valentino *et al.* 2011). There is direct evidence that Arab parties in the Israeli parliament have used the imposition of summoned committees as a mobilizing argument during national elections.<sup>14</sup> Using such tactics is expected, given the work in psychology that has suggested that the perception of threat motivates affiliation with others who also feel threatened (Gump and Kulik 1997; Lowenstein *et al.* 2001; Schachter 1959).

The effect of ethnically biased policy change on EPE is straightforward as it reduces the perceived responsiveness of the government in the eye of the

<sup>&</sup>lt;sup>13</sup> See Ben-Bassat and Dahan (2012) for a review.

<sup>&</sup>lt;sup>14</sup>See, for example, the following video by RAAM TAAL (an Arab party in the Israeli parliament) on youtube: https://www.youtube.com/watch?v=JEzDRjk9RQw

minority group, especially among the target population. The effect of a reduced EPE on political participation is supposed to be quite small given the already low degree of trust in the central government among Arabs as cited above. Thus, the mentioned policy change is predicted to result in lower trust in government responsiveness together with a higher level of IPE among individuals of the target population affiliated with the minority group which resembles the mistrust-low sense of political efficacy hypothesis suggested first by Gamson (1968). The lower level of EPE creates the need to act (demand side) and the higher sense of internal efficacy produces the ability to act (supply side). Arab residents in municipalities with an appointed committee are expected to experience a higher degree of IPE and as a result are anticipated to be more susceptible to political mobilization as compared to general Arab population in municipalities without intervention (or before the intervention). After all, the central government appointees are present in their own municipality. Therefore, Arab individuals who live in intervened municipalities (target population) are predicted to have higher rise in political participation than the rest of the Arab population in Israel, which benefits the Arab parties who gain more votes.

As Beeri and Yuval (2013) found, the target population in Jewish municipalities opposed the intervention which implies that residents in intervened municipalities have perceived the removal of their local elected representatives as yet another way to cut poor sections in the Israeli society from participating in local decision making rather than an insurance device to deal with failed mayors (electoral-error-correction). The intervention may even further decrease the low sense of EPE that characterizes poor people in general. Unlike the Arab citizens, no organization or political party has emerged to frame the policy intervention in terms of social affiliation to mobilize voters in Jewish citizens in affected municipalities. As a result, the policy intervention should dampen their perceived ability to affect the political process and the subsequent policy. The likely response to such intervention in Jewish municipalities is a reduction in both the level of EPE and their sense of IPE that contributes to a lower level of political participation. By contrast, individuals in the Jewish general population (i.e., Jewish municipalities without an appointed council) are likely to gain from the change in public policy. The intervention would reduce the potential bailout costs associated with poor management or strategic behavior by municipalities in weak fiscal conditions. Thus they are anticipated to back the intervention and, indeed, Beeri and Yuval lend support to that prediction. Serving the interests of residents in municipalities with healthy fiscal conditions should increase their sense of EPE as well as IPE and as a result their political participation.

We assume that the Ministry of the Interior (MOI) has not deliberately targeted Arab municipalities to impact their voting behavior. The MOI has decided to appoint summoned committees in certain municipalities according to pre-specified rules and usually after employing less intrusive measures such as rescue plans and appointment of an accountant that reports directly to the central government. Moreover, such a decision has to be made after a recommendation of a professional committee, concluding that a certain municipality is not functioning. A bureaucratic hearing process is started after the Minister of the Interior first states his intention to appoint a summoned committee in a particular municipality. The elected officials of that municipality are provided with the opportunity to convince the regional councilor and then the MOI general director to avoid such a choice. Moreover, the decision is subject to a judicial review and many elected mayors in fact file an appeal to the Supreme Court on the MOI decision to appoint a summoned committee in their municipalities. In Israel, the Supreme Court has the authority to overturn such decisions, but none of the appointed summoned committees were overruled. It should be seen in light of the well-known reputation of an activist Supreme Court that is sensitive to discriminatory policy against Arab citizens in Israel.

Our assumption does not rule out the possibility that the MOI made an explicit decision to trade the financial gains of such a decision for the political costs of an increase in voter turnout in general elections and stronger support for Arab parties. However, it seems that this possibility is less plausible after a conversion with key players. In March 2018, we interviewed separately the Minister of the Interior (Roni Bar-On) and his Director General (Ram Belnikov), who were in charge at the peak of appointing 15 summoned committees during their short term in the MOI (May 2006-July 2007). Both of them emphasized during the interview that the appointment of summoned committees was made after a very careful and meticulous decision process because of the negative democratic consequences of a decision to appoint a summoned committee. While they were aware of the expected resentment, they explicitly have stated on record that the potential political consequences were not part of the decision process in the MOI. We believe their answers are reliable especially given that both of them had retired from political life and public service. Their account is in line also with Ben-Bassat, Dahan, and Klor (2016b) who have found that municipalities were chosen by the MOI based on fiscal measures such as debt, deficit, and taxes. Moreover, assuming that the MOI selected the municipalities that are not expected to react with higher turnout and more votes to Arab parties as a result of appointed committees should drive the relation between voting turnouts and appointed committees toward zero and against our main hypothesis. Thus, for our theoretical framework, it would be plausible to assume that municipalities were selected by the MOI based on their fiscal stress measures.

We believe a convincing case has been established that the intervention in the form of removal of local elected officials has generated an impact on both IPE through its effect on perceived discrimination and EPE; all the while, we should admit that the empirical examination represents a reduced-form estimation of the link between policy change and political participation fluctuations, assuming that IPE and EPE are the mediating factors. We summarize this discussion by presenting our four key hypotheses:

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Hypothesis 1: Minority (perceived) biased public policy should raise IPE and lower EPE among individuals who are affiliated with that minority group, which may result in higher political participation. In our case, the imposition of an appointed committee in Arab municipalities is expected to increase the political participation of their residents in national elections as compared to Arab residents in municipalities without appointed committees.

Hypothesis 2: Members of a majority group who are affected negatively by a change in public policy are expected to show a lower level of both IPE and EPE, which leads to subdued voting turnout. In our case, the imposition of an appointed committee in Jewish municipalities is expected to reduce the political participation of their residents in national elections in comparison with Jewish residents in municipalities without appointed committees.

Hypothesis 3: Arab residents in municipalities where elected officials were replaced by summoned committees would increase their votes for Arab parties in national elections as compared to Arab residents in municipalities without appointed committees.

Hypothesis 4: Jewish residents in municipalities would not change their vote choice regarding the share of votes for Arab parties in national elections following the imposition of an appointed committee.

#### **Empirical Analysis**

#### **Data and Methods**

To test our hypotheses, we exploit a panel dataset on actual voting turnout and political preferences in the last three national elections in Israel for all municipalities. Fixed effects models are used to estimate the impact of perceived biased policy on political behavior. Such models compare between municipalities with and without a summoned committee neutralizing municipalities' unobserved constant characteristics which is necessary to identify the isolated impact of the examined policy on voter turnout and political preferences. Note that, according this research design, the outcome variables (voter turnout and political preferences) in municipalities before and after a summoned committee was appointed are compared to the evolution of the same outcome variables in municipalities that have not experienced a policy change in the same time period. Employing fixed effects models dictate a separate estimation for Arab and Jewish municipalities. One may experiment with one sample with interaction term for Arab municipalities with an appointed committee rather than two separate models for the two separate subsamples of Arab and Jewish municipalities. However, such an estimation method offers an inferior identification strategy because fixed effects should be omitted and therefore unobserved municipalities' characteristics may bias the estimated impact.

We offer a fixed effects model with time-varying municipalities' characteristics in addition to a baseline fixed effects model that includes policy intervention only given that the intervention is not random but rather those municipalities were chosen by the MOI based on their fiscal indicators. Furthermore, these variables may respond endogenously to an intervention (e.g., population size may increase or decrease as a result of an intervention) and impact political behavior. Omitting the time-varying control variables may thus bias the coefficients of interest. In the empirical analysis, we address this concern by accounting for both fiscal measures such as debt, deficit, and taxes as well as outcome variables like income per capita and education achievements. In addition, we plan to employ a series of sensitivity analyses to check the robustness of our results. For example, we will examine whether the results are sensitive to the duration of intervention, softer forms of interventions such as removing the council members only and an appointed accountant.

Unlike most studies, the risk of reverse causality in our research design is rather low. It is unlikely that the removal of elected mayors is driven by political participation and vote choice. Nevertheless, we plan to explore the possibility of reverse causality by examining the relations between past voting turnout in national elections and the likelihood of summoned committee.

We summarize the above methodological discussion by the following statistical model:

$$Turnout_{i,t} = a(appointed \ committee)_{i,t} + \mathbf{X}_{i,t}\mathbf{b} + c_t + d_i + u_{i,t}$$
(1)

where  $Turnout_{i,t}$  is the share of eligible voters who show up in the ballot boxes in municipality *i* at election *t*. The summoned committee is a dummy variable that gets 1 if the authority in municipality *i* is in the hands of appointed officials at the time of election *t*, and zero otherwise. We use the same timing for both turnout and summoned committee to capture the saliency of that central intervention during the national election.  $\mathbf{X}_{i,t}$  is a vector of time-varying municipality characteristics;  $c_t$  is a fixed effect for each election year in the sample to control for general trends;  $d_i$  is a municipality fixed effect; *a* and the vector **b** are unknown parameters that would be estimated. Unobserved determinants of political participation at the municipality level are represented by the error term,  $u_{i,t}$ .<sup>15</sup> We estimate equation 1 for Arab and Jewish municipalities separately, and as suggested in Hypotheses 1 and 2, the estimated coefficient *a* should be positive in Arab localities and negative in Jewish municipalities.

<sup>&</sup>lt;sup>15</sup> The standard errors are clustered at the municipality level.

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To examine the effect of a policy change or perceived discrimination on vote choice, we estimate the following regression:

Arab parties<sub>*i*,*t*</sub> = 
$$\alpha$$
(appointed committee)<sub>*i*,*t*</sub> +  $Z_{i,t}\beta$  +  $\gamma_t$  +  $\delta_i$  +  $e_{i,t}$  (2)

where the dependent variable is measured by the share of votes (out of total votes) for the three main Arab parties in the national parliament (KNESET) in municipality *i* at election year *t*. As before, the appointed committee is a dummy variable that gets 1 if the authority in municipality *i* is in the hands of appointed officials at the time of election *t*, and zero otherwise.  $Z_{i,t}$  is a vector of time-vary-ing municipality characteristics;  $\gamma_t$  is a fixed effect for each election year in the sample to control for general trends;  $\delta_i$  is a municipality fixed effect;  $\alpha$  and the vector  $\beta$  are unknown parameters that would be estimated. Unobserved determinants of vote choice at the municipality level are represented by the error term,  $e_{i,t}$ .<sup>16</sup> We estimate the above model for Arab and Jewish municipalities separately, and as suggested in Hypothesis 3, the estimated coefficient  $\alpha$  should be positive in Arab localities.

Estimating equation 2 should be seen as a test of the suggested mechanism behind political participation. The Arab parties are key elite agents of Arab citizens in Israel who frame the mentioned policy change as an act of discrimination and provide organized resources to mobilize Arab voters to cast their protest in the ballot box. Thus perceived discrimination that is associated with a policy change increases political engagement of Arab citizens in national elections. Estimating the impact of that policy change on vote choice of Jewish citizens could be seen as a placebo test. As stated in Hypothesis 4, the expected coefficient  $\alpha$  should be zero in Jewish municipalities.

#### Data

The empirical analysis covers all municipalities that have data on all the variables that are used in this study. Our research is therefore based on 188 local municipalities in 2006 and 2009 elections and 195 in the 2013 election out of 201, which is the total number of municipalities in Israel (see Table 1). Regional municipalities were excluded due to lack of data on voter turnout and vote choice.<sup>17</sup>

<sup>&</sup>lt;sup>16</sup> As before, we cluster the standard errors at the municipality level.

<sup>&</sup>lt;sup>17</sup> Five Druz municipalities in the Golan Heights (Buq'ata, Majdal-Shams, Mas'ade, Ein-Qinyye, and Ghajar) as well as the municipality of Basma (an Arab municipality in the Galilei) are excluded in all three elections due to lack of data on voter turnout. In addition, seven other municipalities are excluded to avoid potential bias due to the effect of merged municipalities. These seven localities were merged in 2003 and a few years later were unmerged following considerable pressure by their residents (the seven municipalities are Ba'ana, Dir-el-Asad, and Majd-el-Crum in 2008, Daliet-el-Carmel and Usefia in 2008, and Baqa-Al-Gharbiyye and Jat in 2010).

	All Municipalities	1	ities without 10n Board	1	lities with a on Board
		Jews	Arabs	Jews	Arabs
2006	188	119	66	1	2
				(.39)	(.38)
2009	188	113	56	7	12
				(2.10)	(1.80)
2013	195	114	58	6	17
				(5.37)	(4.57)
Total	571	346	180	14	31
				(3.38)	(3.23)

#### Table 1. Number of Municipalities and Intervention Duration, by Election and Sector

*Source*: Data provided by Michal Goldstein, Senior Director of the Department of Municipal Administration, at the Local Government Administration, Ministry of Interior and the manuscript: "The duration of Summon board," Research and Information Center, the Knesset, 2009.

*Note*: In parentheses are the number of years the summon board committee serves till the date of next K nesset elections.

We have merged data from two sources—the Interior Ministry and the Center for Research and Information at the Knesset—to generate time series data on the timing and duration of appointing a summoned committee and an accountant in the years 2005-13. This is our key independent variable. The total number of summoned committees that were appointed is 45 and the duration of intervention is more than three years on average during the period covered in this research (see Table 1). Table 2 presents a detailed timeline of the appointed committees for each municipality.

The data on actual average voter turnout (actual voters relative to eligible voters) in national election and the average voting share for Arab political parties (out of the total votes) in each municipality for the last three national elections in the years 2006, 2009, and 2013 is taken from the official Knesset website. This is also the source of data on other dependent variables such as the share of votes for the party of the Interior Minister, Prime Minister, or Coalition Parties. As can be seen in Table 3, the average voter turnout in national elections is around 62 percent (this is a simple average across municipalities without attaching weights to each municipality according to its size in terms of population). The political participation in Jewish municipalities is higher as compared to Arab municipalities by nine percentage points. A large share of Arab citizens votes for Arab parties while they gain only very negligible political support in Jewish municipalities (see Table 3).

The Central Bureau of Statistics is the source of time-varying municipality characteristics such as the size of population in each municipality, demographic

		L	Table 2. Municipalities with a Summon Board, 2005-13	cipalities wi	th a Summe	on Board, 20	05-13			
	2004	2005	2006	2007	2008	2009	2010	2011	2012	2013
Municipality			28/3/06			10/2/09				22/1/13
Abu Ghosh				~	>					
Ofakim				~	>	$\mathbf{r}$	~	$\mathbf{i}$	>	$\overline{}$
Ibillin		~	7	~	>	~	~			
Beit Jann		~	7	~	>					
Jisr al –Zarqa								$\mathbf{i}$	>	$\overline{}$
Daburiya				~	>	$\mathbf{r}$	~	$\mathbf{i}$	>	~
Zemer				~	>	$\mathbf{r}$	~	$\mathbf{i}$	>	$\overline{}$
Zarzir							~	$\mathbf{i}$	>	~
Hadera							~	$\mathbf{i}$	>	$\overline{}$
Tuba Zangariyye				~	>	~	~	$\mathbf{i}$	>	$\overline{\mathbf{v}}$
Taybeh				~	>	$\mathbf{h}$	~	$\mathbf{i}$	>	$\overline{}$
Tira				~	>					
Tamra						1/12/09	~	$\mathbf{i}$	>	$\overline{}$
Turan				~	>	$\mathbf{r}$	~	$\mathbf{i}$	>	$\mathbf{a}$
Yanuh-Jat						1/12/09	~	$\mathbf{i}$	~	$\overline{}$
Yesod Hamala				~	>	$\checkmark$	~	$\mathbf{i}$	>	$\overline{}$
Yeruham		~	7	7	$\mathbf{r}$	$\checkmark$	7			
Yarqa				~	$\mathbf{r}$	$\checkmark$	~	$\mathbf{i}$	$\mathbf{r}$	$\overline{}$
Kefar Kana				~	$\mathbf{r}$	$\mathbf{r}$	~	$\mathbf{i}$	$\mathbf{r}$	$\mathbf{k}$
Kafr Manda				~	>	$\checkmark$	~	$\mathbf{i}$	~	$\overline{}$
Lod				~	$\mathbf{r}$	$\checkmark$	~	$\mathbf{i}$	$\mathbf{r}$	~
Migdal				$\checkmark$	$\checkmark$	γ	$\checkmark$	V	$\checkmark$	$\checkmark$
										Continued.

				Table 2. (	Table 2. (Continued)					
	2004	2005	2006	2007	2008	2009	2010	2011	2012	2013
Miilya				~	~			~	$\mathbf{r}$	~
Maale Iron			1/11/06	$\overline{}$	~	~	$\mathbf{r}$	~	~	~
Mitzpe Ramon				$\mathbf{i}$	~					
Nahf			27/9/06	$\mathbf{i}$	~	~	$\mathbf{i}$	$\mathbf{k}$	$\mathbf{r}$	~
Arad				$\mathbf{i}$	~	7	$\mathbf{k}$			
Arara						1/12/09	~	$\overline{}$	$\mathbf{r}$	$\mathbf{k}$
Arara Banegev					~	~	~	$\overline{}$	$\mathbf{r}$	~
Ktzir Harish			1/10/06	$\mathbf{i}$	~	~	~	$\mathbf{k}$	~	7
Total	0	3	9	23	24	22	24	23	23	23
Total, at the time of elections	e 0	S	б	23	24	19	24	23	23	23
<i>Source</i> : See Table 1 <i>Note</i> : Merged muni	<i>Source</i> : See Table 1. <i>Note</i> : Merged municipalities which were later separated were excluded from the statistical analysis.	1 were late	sr separated wer	e excluded f	rom the sta	tistical analysis.				
	Tabi	le 3. Vote	Table 3. Voter Turnout and Voting Share for Arab Parties in National Elections	Voting Sha	re for Arab	Parties in Na	tional El	ections		
		Vote	Voter Turnout (%)			Λ	oting Sha	Voting Share for Arab Parties (%)	rties (%)	
•	All Municipalities	Jewish	Jewish Municipalities	Arab Municipalities	icipalities	All Municipalities	ities	Jewish Municipalities		Arab Municipalities
2006	62.1 (11.1)	9	64.0 (9.2)	58.6 (13.1)	[13.1]	24.4 (34.9)		.8 (3.1)	6:	65.9 (25.5)
2009	61.0 (12.3)	9	65.3 (8.5)	53.5 (14.2)	14.2)	28.0 (39.7)		1.0 (3.2)	7.	75.7 (27.8)
2013	64.0 (11.4)	9	68.2 (9.1)	57.3 (11.6)	11.6)	27.6 (37.8)		1.0 (3.5)	Ż	70.1 (27.4)
Average	62.4		65.8	56.5	.5	26.7		1.0		70.5
<i>Notes</i> : The numb ages across mun in parentheses.	<i>Notes</i> : The number of municipalities is 188 in 2006 and 2009 elections and 195 in 2013 election. The voter turnout and voting share are simple aver- ages across municipalities (i.e., the weight attached to each municipality is not influenced by its size in terms of population). Standard errors are in parentheses.	es is 188 ir e weight at	n 2006 and 2009 ttached to each	elections an municipality	nd 195 in 20 y is not infl	l3 election. The uenced by its siz	voter tur. ze in tern	nout and votin <sub>.</sub> as of populatio	g share are n). Standa	e simple aver- urd errors are

composition (the ratio of the population at the age of 65 or over to 20-64), average monthly wage, and education level (the share of eligible students for matriculation in 12th grade). The socioeconomic variables at the municipality level are computed for election years except for the last election for which we use 2012, that is the most recent available data.

#### Results

It is worth illustrating the central results by looking at simple diff-in-diff tabulation. In Table 4, we compute simple averages of our two key dependent variables for two groups of municipalities: municipalities with an appointed committee either in 2009 or 2013 and municipalities without intervention. As can be seen, the voter turnout in national elections "before" the intervention in Arab municipalities is statistically the same in both groups of municipalities regardless of intervention. Comparing before and after (first diff) and between intervened and nonintervened Arab municipalities (second diff), we find that the intervention raises significantly voter turnout in national elections. The lower panel of Table 4 presents a significant increase also in voting share for Arab parties in Israeli parliament.

Table 5a shows the first main finding of this study using ordinary least squares regressions. In line with Hypothesis 1, the effect of a summoned committee is negative on political participation in Jewish municipalities and positive in Arab municipalities. The effects are significant and relatively large in Arab municipalities. The removal of an elected mayor who is replaced by an appointed committee tends to raise voter turnout by approximately five percentage points in Arab municipalities using the estimated coefficient. This suggests that our proxy (an appointed committee) for perceived group discrimination does mobilize Arab citizens to show up at the ballot boxes in larger numbers. As can be seen in Table 5b, Arab parties benefit from the higher voter turnout in national elections. The share of votes for Arab parties is higher by 4.4 percentage points but this coefficient is borderline significant. The estimated coefficient is quite large especially given that the vote share of Arab parties is already around 70 percent (see Table 3). As expected, the "treatment" has no significant impact in terms of vote choice in Jewish towns and cities, which should be seen also as a placebo test. We have estimated the effect of a summoned committee on vote choice also in terms of vote share for the party of the Interior Ministry, Prime Minister, and coalition parties and none come up consistently significant (the results are not reported here).

One may think that the results may reflect a compensation effect whereby those residents who are not allowed to cast their vote at local elections compensate themselves in national elections. The compensation hypothesis implies the same predicted effect in both Arabs and Jews—which is not consistent with our results. Moreover, Ben-Bassat and Dahan (2012) show that this compensation

				Table 4. III	Table 4. Illustrating the Effects: Diff-in-Diff	e Effects: L	)iff-in-Diff				
		2006			2009			2013		Diff-in-diff	n-diff
.1	Non- intervened	Intervened	Intervened Difference	Non- intervened	Intervened	Intervened Difference	Non- intervened	Intervened	Intervened Difference	2006/2009	2006/2013
Voter turnout											
Arab	.594 (.019)	.563 (.029)	031 (.035)	.529 (.021)	.554 (.030)	.026 (.037)	.570 (.017)	.581 (.025)	.011 (.030)	.057*** (.021)	.042** (.021)
Jewish	.645	.581	063***	.657	.595	062***	.687	.618	068***	.001	005
	(600.)	(.022)	(.024)	(800.)	(.019)	(.021)	(600.)	(.022)	(.023)	(.007)	(.007)
All	.629	.568	061***	.617	.567	050**	.651	.593	058***	.010	.003
	(600.)	(.021)	(.023)	(.010)	(.022)	(.024)	(600.)	(.018)	(.020)	(.013)	(.013)
Vote share for Arab parties	Arab part.	ies									
Arab	.663	.648	015	.743	.794	.051	.701	.732	.032	.066**	.047*
	(.036)	(.062)	(.072)	(.040)	(.064)	(.075)	(.040)	(.054)	(.068)	(.030)	(.028)
Jewish	.008	.019	.011	600.	.021	.012	600.	.024	.015	.001	.003
	(.003)	(.014)	(.014)	(.003)	(.015)	(.015)	(.003)	(.018)	(010)	(.001)	(:005)
All	.210	.454	.244***	.236	.557	.321***	.223	.514	.292***	$.076^{***}$	.047**
	(.026)	(.072)	(.076)	(.030)	(.083)	(.088)	(.028)	(.075)	(000)	(.023)	(019)
<i>Notes</i> : Municipalities with an appointed committee either in 2009 or 2013 are defined as "intervened." According to that definition, there are 18 Arab municipalities with intervention and 50 without. There are eight Jewish municipalities with intervention and 112 without. ************************************	ipalities wi alities with level of 10%	th an appoi 1 interventic 6, **signific	inted commit on and 50 wit ance level of	tee either ii hout. There 5%, ***signi	n 2009 or 201 e are eight Jer ficance level	13 are define wish municij of 1%. Clust	d as "interv palities with tered stands	ened." Acco intervention ind errors are	rding to tha 1 and 112 wi 2 in parenthe	t definition, thout. sses.	there are 18

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Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.016 (.011)	010** (.004)	.049*** (.016)
Adjusted R-squared	.15	.65	.21
Observations	571	360	211
Observations with a summon board	45	14	31
Number of municipalities	195	120	75

#### Table 5a. The Effect of a Summon Board on Voter Turnout in National Elections

*Notes*: \*Regressions were estimated using ordinary least squares with municipality and election year fixed effects. \*Significance level of 10%, \*\*\*significance level of 5%, \*\*\*significance level of 1%. Clustered standard errors are in parentheses.

A merger of several municipalities is considered as a unified municipality after the time of the merger. Merged municipalities which later were separated were excluded.

Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.041** (.017)	.001 (.003)	.044* (.022)
Adjusted R-squared	.15	.05	.34
Observations	571	360	211
Observations with a summon board	45	14	31
Number of municipalities	195	120	75
Note: See Table 5a.			

#### Table 5b. The Effect of a Summon Board on Voting Share for Arabs Parties

hypothesis, which is very common with regard to Arab citizens, is unsupported empirically. In fact, they found that citizens in Arab municipalities who tend to vote at higher rates in local elections are also more likely to show up in the ballot boxes in national elections.

We deal here with the risk of endogeneity that might stem from the fact that the MOI has selected those municipalities based on their fiscal stress indicators which may be correlated with municipality characteristics. Indeed, the municipalities with summoned boards tend to be poorer (a lower average income) and smaller (a lower population size). As a result, the estimated effect may capture the differences between municipalities that were selected by central government rather than what the intervention does to political behavior. One more channel through which municipality characteristics may bias our results is their potential dynamics impact on attitudes. We address this risk by controlling for timevarying municipality characteristics such as socioeconomic status, education level, population size, and its composition, in addition to municipality fixed effects. Accounting for these municipality features also deals with the potential indirect effect that intervention may have on political behavior through its impact on individuals' resources like average income, as suggested by Verba, Schlozman, and Brady (1995).<sup>18</sup> Moreover, in our empirical setting the control group includes municipalities that have received "treatment" in previous or subsequent national elections which makes them more comparable to "treated" municipalities. As can be seen from Table 2, no municipality had an appointed committee in all three national elections that are covered in our study: nine municipalities were ruled by an appointed committee during one national election only and 18 municipalities had such committees during two elections.

As can be seen in Tables 6a and 6b, all four hypotheses that are sketched above gain support after controlling for time-varying municipality characteristics.<sup>19</sup> To make the comparison more meaningful, the first panel of Table 6a presents the estimated coefficient with observations that have complete data on municipality characteristics. The coefficient of an appointed committee is .036 with time-varying municipality characteristics and .042 without them. However, the effect of intervention on the vote share for Arab parties is smaller and less precisely estimated.

Nevertheless, we could not exclude the risk of selection on unobservable municipality characteristics. However, the selection of municipalities should be uncorrelated with the mentioned list of control variables including socioeconomic status, education level, population size, and its composition to bias the results. The amount of selection on unobservable to selection on the observed explanatory variables would have to be high to mistrust our estimates (Altonji, Elder, and Taber 2005). Our results gain more confidence given that the estimated coefficient remains almost the same as in Table 5a.

In Tables 7a and 7b, we control for municipality fiscal distress indicators (taken from Israeli CBS).<sup>20</sup> These include municipality debt relative to revenues and tax effectiveness measures (collected local taxes relative to local taxes that should have been collected) which were computed a year before election (the results are similar adding budget deficit relative to revenues but are not reported here). Note that, according to the stated rules of the MOI, the appointed committee is imposed in municipalities that cross certain thresholds of these three fiscal indicators. This list of additional control variables aims to capture the potential dynamic in residents' views regarding the intervention. Residents in

<sup>20</sup> Israel Central Bureau of Statistics, Local authorities in Israel (various years).

<sup>&</sup>lt;sup>18</sup> Afriat and Dahan (2010) show that political participation in national elections in Israel is positively affected by the level of earnings.

<sup>&</sup>lt;sup>19</sup> Using the actual net immigration to a certain municipality, which is a common measure of municipality attractiveness in Israel, instead of population size generates the same results (not reported here).

	Table 6a. The Ef	fect of a Summon	Board on Voter Tu	Table 6a. The Effect of a Summon Board on Voter Turnout in National Elections	llections	
		Sensitivity Analy	sis: Adding Municij	Sensitivity Analysis: Adding Municipality Socioeconomic Characteristics	c Characteristics	
		(1)			(2)	
Variable	All Municipalities	Jews Municipalities	Arabs Municipalities	All Municipalities	Jews Municipalities	Arabs Municipalities
A summon board served at the time of Knesset election	.013 (.012)	007 (.004)	.042** (.016)	.010 (.011)	008	.036** (.015)
The ratio of 65+ to 20-64				.043	026	014
population				(.149)	(.078) 077***	(11) 018
Log population				006 (.054)	.016)	018 (.145)
Log average monthly wage				.202***	.092***	.386***
of employees				(.071)	(.029)	(.135)
Rate of eligibility for				002	002	045
matriculation among 12th grade students				(.044)	(.022)	(090)
Adjusted R-squared	.13	.65	.17	.16	.71	.22
Observations	517	323	194	517	323	194
Observations with a summon board	38	6	29	38	6	29
Number of municipalities included in the regression	184	111	73	184	111	73
Note: See Table 5a.						

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		Sensitivity Analys	Sensitivity Analysis: Adding Municipality Socioeconomic Characteristics	ality Socioeconomic	Characteristics	
		(1)			(2)	
Variable	All Municipalities	Jews Municipalities	Arabs Municipalities	All Municipalities	Jews Municipalities	Arabs Municipalities
A summon board served at the time of Knesset election	.043*** (.013)	.003 (.004)	.039** (.016)	.037*** (.012)	.003 (.004)	.029* (.015)
The ratio of 65+ to 20-64				.339**	025	.763
population				(.153)	(.025)	(.730)
Log population				.023	008	760.
				(.044)	(900)	(.150)
Log average monthly wage				.149**	005	.354**
of employees				(.067)	(900)	(.152)
Rate of eligibility for				062	$010^{**}$	.003
matriculation among 12th grade students				(.046)	(.005)	(.073)
Adjusted R-squared	.16	.06	.33	.19	60.	.37
Observations	517	323	194	517	323	194
Observations with a summon board	38	6	29	38	6	29
Number of municipalities included in the regression	184	III	73	184	111	73
Note: See Table 5a.						

	Table 7a. T	Table 7a. The Effect of a Summon Board on Voter Turnout in National Elections	non Board on Voter	Turnout in Nation	al Elections	
		Sensitivit	y Analysis: Adding I	Sensitivity Analysis: Adding Municipality Fiscal Indicators	ndicators	
		(1)			(2)	
Variable	All Municipalities	Jews Municipalities	Arabs Municipalities	All Municipalities	Jews Municipalities	Arabs Municipalities
A summon board served at the time of Knesset election	.019 (.012)	009** (.004)	.055*** (.017)	.022* (.012)	014*** (.005)	.053*** (.017)
Debt (as a share of revenues)				008	.021**	007
Budget deficit (as a share of				(010.)	(010.)	(170.)
Tax effectiveness				013	.011	.018
Adjusted R-squared	.15	.66	.21	(100.)	(070.)	(oco.) 12.
Observations	527	334	193	527	334	193
Observations with a summon board	42	13	29	42	13	29
Number of municipalities included in the regression	192	119	73	192	119	73
Note: See Table 5a.						

	Table 7b	Table 7b. The Effect of a Summon Board on Voting Share for Arab Parties	mmon Board on Vo	ting Share for Arat	Parties	
		Sensitivit	y Analysis: Adding N	Sensitivity Analysis: Adding Municipality Fiscal Indicators	ıdicators	
		(1)			(2)	
Variable	All Municipalities	Jews Municipalities	Arabs Municipalities	All Municipalities	Jews Municipalities	Arabs Municipalities
A summon board served at the time of Knesset election	.042** (.018)	.002 (.003)	.047* (.024)	.042** (.024)	.002 (.003)	.050* (.025)
Debt (as a share of				008	.005	004
revenues)				(.027)	(.007)	(.036)
Budget deficit (as a				.030	011	.049
share of revenues)				(.035)	(.013)	(.036)
Tax effectiveness				019	006	049
				(.050)	(.004)	(.065)
Adjusted R-squared	.15	.06	.33	.15	.06	.33
Observations	520	329	191	520	329	191
Observations with a summon board	41	13	28	41	13	28
Number of municipalities included in the regression	161	118	73	191	118	73
Note: See Table 5a.						

municipalities with an appointed committee may first resent central government intervention, but might change their minds based on actual fiscal outcomes. The attitudes toward central government intervention could be mitigated as a result of changes to the well-being of the residents in affected municipalities that may be attributed to the intervention. Table 7a shows that the "treatment" effect is significant and quantitatively important but somewhat smaller (.055 as compared to .053) when fiscal indicators are not controlled for. The estimated coefficient of intervention on Arab parties vote share remains almost the same (Table 7b).

Another seemingly source of bias could be the signal that the intervention itself may send to individuals regarding the economic distress of their municipality, that might affect negatively, for example, their expected permanent income. According to this suggestion, the change in political behavior may be the result of that signal rather than how public policy is perceived by the residents in intervened municipalities. Our empirical design addresses this concern by estimating the above model for Arab and Jewish municipalities separately to allow for an interaction effect between dominant/minority affiliation and perceived discrimination due to a particular policy change. The signaling story should have the same effect in both Arab and Jewish municipalities. By contrast, we found a positive coefficient in Arab localities but a negative one in Jewish municipalities, which is consistent with our Hypotheses 1 and 2. Furthermore, our empirical design is immune to a general policy change that may affect uniformly the political behavior of all Arab citizens due to hostile political climate or following wide-ranging discriminatory policy like the proposal to downgrade Arabic from being an official language to a lower status. Such a policy should evenly impact political participation in both types of Arab municipalities, which is inconsistent with our hypotheses and findings.

Our main results are not affected considerably after adding the duration of the appointed committee to our baseline regressions, except that the estimated coefficient has a higher standard deviation (see Tables 8a and 8b). We do not control for socioeconomic characteristics in these regressions (and in the next regressions) to keep the number of "treated" observations as large as possible. Adding socioeconomic variables to the robustness test generates the same general picture, but with less precision and significance.

The second robustness test is to see whether the results are sensitive to the inclusion of softer central government interventions such as appointing an accountant. The estimated coefficient of our key explanatory variable in Tables 9a and 9b remains significant and with almost the same magnitude after adding an appointed accountant to our baseline regression. An appointed accountant has no significant effect on voter turnout in national elections or vote share for Arab parties. The lower saliency of this sort of intervention, as compared to the high visibility of the removal of elected mayors, may explain the zero effect we find.

	5	ysis: Adding the D non Board Commit	
Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.033** (.016)	008 (.009)	.066*** (.022)
Duration of summon board committee service (month)	000* (.000)	000 (.000)	000 (.000)
Adjusted R-squared	.15	.65	.21
Observations	571	360	211
Observations with a summon board	45	14	31
Number of municipalities	195	120	75
Note: See Table 5a.			

#### Table 8a. The Effect of a Summon Board on Voter Turnout in National Elections

#### Table 8b. The Effect of a Summon Board on Voting Share for Arab Parties

Sensitivity Analysis: Adding the Duration of the Summon Board Committee

Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.052** (.026)	002 (.002)	.052 (.033)
Duration of summon board committee (month)	000 (.000)	.000 (.000)	000 (.001)
Adjusted R-squared	.15	.06	.34
Observations	571	360	211
Observations with a summon board	45	14	31
Number of municipalities	195	120	75
Note: See Table 5a.			

Next, we limited our sample to the cases where both the elected mayor and the council members were removed (i.e., the dismissal of the elected council only has not been considered as an intervention or "treatment"). As before, the general picture that emerges is quite similar (see Tables 10a and 10b), except that the estimated effect on Arab parties is less precise.

	Sensitivity	Analysis: Adding a Accountant	an Appointed
Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.017 (.012)	011*** (.004)	.049*** (.016)
With an appointed accountant at the time of Knesset election	011 (.008)	009** (.004)	.001 (.015)
Adjusted R-squared	.15	.65	.20
Observations	571	360	211
Observations with a summon board	45	14	31
Observations with an appointed accountant	185	47	138
Observations with a summon board and an appointed accountant	34	8	26
Number of municipalities	195	120	75
Note: See Table 5a.			

## Table 9a. The Effect of a Summon Board on Voter Turnout in National Elections

### Table 9b. The Effect of a Summon Board on Voting Share for Arab Parties

Variable	Sensitivity Analysis: Adding an Appointed Accountant		
	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.040** (.017)	.001 (.003)	.042* (.021)
With an appointed accountant at the time of Knesset election	.021 (.013)	001 (.002)	.017 (.020)
Adjusted R-squared	.16	.05	.35
Observations	571	360	211
Observations with a summon board	45	14	31
Observations with an appointed accountant	185	47	138
Observations with a summon board and an appointed accountant	34	8	26
Number of municipalities	195	120	75
Note: See Table 5a.			

Variable	Members Had Been Removed		
	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.012 (.012)	011*** (.004)	.052** (.021)
Adjusted R-squared	.14	.65	.19
Observations	571	360	211
Observations with a summon board	36	13	23
Number of municipalities	195	120	75
Note: See Table 5a.			

#### Table 10a. The Effect of a Summon Board on Voter Turnout in National Elections

#### Table 10b. The Effect of a Summon Board on Voting Share for Arab Parties

Sensitivity Analysis: Both the Mayor and Council Members Had Been Removed

Sensitivity Analysis: Both the Mayor and Council

Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
A summon board served at the time of Knesset election	.046** (.022)	.002 (.003)	.048 (.030)
Adjusted R-squared	.15	.05	.34
Observations	571	360	211
Observations with a summon board	36	13	23
Number of municipalities	195	120	75
Note: See Table 5a.			

The risk that the interior ministry has selected the policy tools (fiscal thresholds) based on political behavior is slim. Nevertheless, one may claim that the central government has targeted Arab municipalities according to their political participation and vote choice trends. According to this (unlikely) story, the true intention of the government in launching this major economic program was curbing the upward trends in voter turnout and vote share for Arab parties in certain Arab municipalities. To address this potential bias that may result from reverse causality, we use the appointment of a summoned committee in 2007 as a dependent variable and the rate of change in voter turnout (or vote share for Arab parties) as the independent variable. The results in Tables 11a and 11b do not lend support to the reverse causality hypothesis.

	Summon Board Served at 2007		
Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
Rate of change in voter turnout in Knesset election (2003-06)	.074 (.310)	452 (.521)	081 (.459)
Constant	.133***	.015	.214***
	(.041)	(.067)	(.063)
Adjusted R-squared	01	0	02
Observations	184	116	68
Note: See Table 5a.			

# Table 11a. Reverse Causality—The Effect of Voter Turnout on a Summon Board in2007

# Table 11b. Reverse Causality—The Effect of Voting Share for Arab Parties on aSummon Board in 2007

	Summon Board Served at 2007		
Variable	All Municipalities	Jewish Municipalities	Arab Municipalities
Rate of change in voting turnout for Arabs parties (2003-06)	010 (.037)	007 (.039)	036 (.066)
Constant	.127*** (.026)	.066*** (.024)	.227*** (.052)
Adjusted R-squared	01	01	01
Observations	174	106	68
Note: See Table 5a.			

#### Conclusion

We used a major policy intervention to study the effect of a biased policy that may convey a negative message concerns the Arab minority group that impacts upon their voting turnout and political preferences. To explore the link between policy and politics, we used the appointment of a summoned committee in a large number of Arab and Jewish municipalities in Israel, which has been perceived as ethnically biased by an Arab minority (Arab citizens). We show that Arab voters in intervened municipalities are more likely to show up at the ballot boxes in national elections. This result is supported by three important complementary findings. First, the effect of policy on political behavior found in Arab municipalities reflects the different reaction of one group of Arab municipalities (i.e., with summoned committees) compared to other group of Arab municipalities (i.e., without summoned committees). Therefore, the long and troubled relationship between Arabs and Jews in Israel should affect all Arab municipalities regardless of the examined intervention. Second, we show that those Arab voters provided subsequently stronger support to Arab parties as would be expected in view of our theoretical framework. Last, the same intervention had an opposite effect in Jewish municipalities. The political participation of Jewish citizens is lower in municipalities with appointed councils without a noticeable effect on vote choice. These important results are obtained using *actual* data on political participation and voting preferences in national elections. We believe this study presents credible evidence that the intervention affected the attitudes of perceived discrimination which, in turn, have impacted on political efficacy. Yet, we do acknowledge that the empirical examination reflects an estimation of a reduced form of relation between a public policy and political behavior, assuming that IPE and EPE are the mediating factors.

The main estimated coefficients remain almost the same after controlling for municipality and election year fixed effects and a long list of time varying municipality characteristics—which may capture both residents' resources and taste, such as socioeconomic status as well as fiscal distress indicators, that may represent changes in residents' attitudes. The robustness of our estimates reduces the risk that the results are driven by unobservable municipality characteristics. These results survive also a series of robustness tests such as adding the duration of intervention and other forms of intervention.

The findings of this study have three important implications. First, it shows that policies have an impact on politics. The likelihood of various social groups to participate in national elections is affected by public policy. In particular, citizens of a certain social group who sense that the central government is implementing policies that discriminate against their political and economic interests chose to vote in larger numbers and vote disproportionally for parties that represent their social group. This result is consistent with mistrust-sense of political efficacy that was first suggested by Gamson (1968). Unlike the previous literature on mass policy feedback, the present article shows that a perceived negative message embodied in public policy may foster or dampen political participation, depending on social group affiliation. The subtle and interesting insight that emerges is that, despite the increased feeling of mistrust among Arab citizens following certain types of public policy, they still play by the rules and believe that conventional political actions are the preferred means to change public policy.

Second, those who are pushing policies that are perceived as antiminority should be aware that they have an impact on their own subsequent political power. We show here that the same policy generates higher political participation of a minority group and reduces the likelihood of members from the majority group to participate in national elections. Moreover, the citizens that are affiliated with the minority group vote more for minority-related parties. Thus the interplay between social structure and the way public policy is perceived may affect the actual makeup of political preferences and therefore the outcome of national elections. The finding that policy actions may stimulate the political engagement of certain groups and attenuate voter turnout of other social groups, thus changing the composition of political preferences expressed at the polls, becomes much more important as nation states continue to diversify in terms of their social structure.

Finally, this article suggests that voter turnout may vary over time even if individual and institutional characteristics remain constant. Minority groups who face considerable adverse policy changes are expected to participate more in elections and strengthen their preexisting voting behavior for a given institutional environment and for the same individual traits. Therefore, explaining the large volatility in political participation and voting behavior over time should take into account the interaction between public policy and social structure.

We believe that our contribution travels beyond the investigated case study of uneven removal of elected mayors in Israel. For example, it is particularly relevant for the current climate in the United States. Specifically, we suggest that the recent immigration rhetoric and the related executive orders by President Trump conveys a negative message which should raise group consciousness among certain social groups such as African Americans, American Muslims, and Latinos. Our theoretical framework predicts that such an ethnically biased policy is expected to lead to a higher voting turnout among minority groups due to a combination of a higher level of IPE and lower EPE. In contrast, voter turnout among members of the majority group should show a slight decline. The results of two recent elections in Alabama and Virginia are in line with our predictions. The high turnout among black voters in Alabama played a decisive role in shaping the outcome of that election. For example, the headline in The Atlantic (December 12, 2017) after special senate elections in Alabama was "African American Voters Made Doug Jones a U.S. Senator in Alabama" (Newkirk 2017). Obviously, our judgment should be seen as suggestive, since the election results are affected by many other factors. The recent state's gubernatorial election in Virginia offers us even more related illustration. The democrat Ralph Northam won that election following a surge in voter turnout among black and Latino voters. A survey conducted by the African American Research Collaborative in partnership with Latino Decisions suggested that racially charged policy platforms by the republican candidate, Ed Gillespie (defending Confederate statues and associating illegal immigration with violent Latinos) played a noticeable role in explaining the strong support among black voters for the democrat candidate (Schneider 2017).

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